

PARAMETER ESTIMATION CONCEPTS

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This is an introduction to the topic.
See book, ~~%~~Parameter Estimation in Engineering and Science+
by J.V. Beck and K.N. Arnold, Wiley, 1977

GENERAL TOPICS:

MEASUREMENT ERRORS

SCALED SENSITIVITY COEFFICIENTS

ESTIMATION METHODS

EXAMPLE

**Parameter estimation is a bridge
between experimentation and analysis/
modeling /computation**

Parameter Estimation vs Function Estimation

Parameter Estimation

Small number of parameters, 2, 3, 4,5

Nonlinear even if ode or pde are linear

Not ill-posed

**Parameters may be a physical property, e.g., thermal
conductivity**

Function Estimation

Large number of parameters, 50, 100 or even 1000

Might be linear if ode or pde are linear

Ill-posed

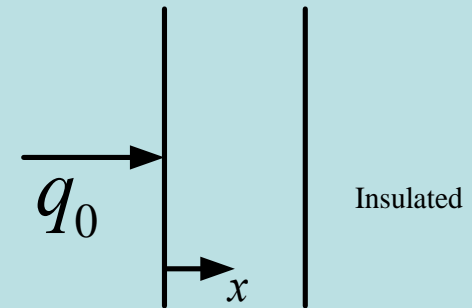
Parameters not a physical property, e.g., heat flux

Heat Conduction Example

$$k \frac{\partial^2 T}{\partial x^2} = C \frac{\partial T}{\partial t}, \quad 0 < x < L, \quad t > 0$$

$$-k \frac{\partial T}{\partial x}(0, t) = q_0, \quad \frac{\partial T}{\partial x}(L, t) = 0$$

$$T(x, 0) = 0$$

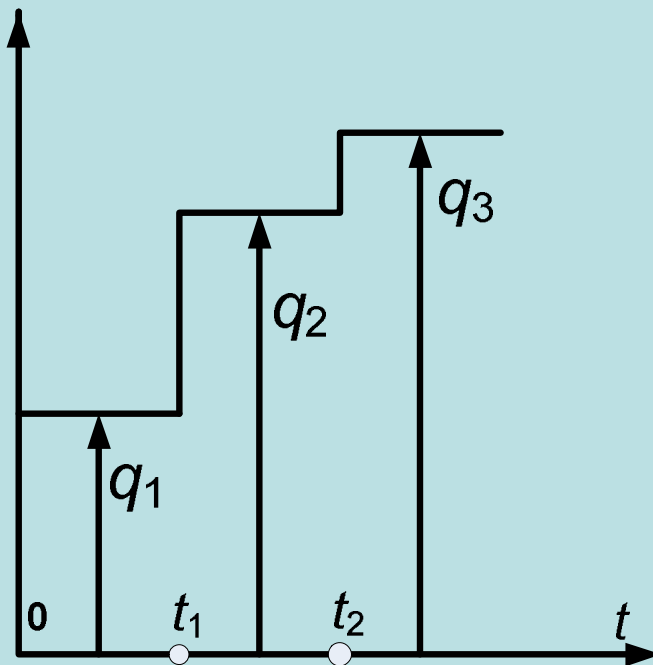


Parameter estimation problem: Estimate k and C given measurements of $T(x, t)$ and q_0

k = thermal conductivity, physical characteristic of a material, $W/(m\ C)$

C = volumetric heat capacity, physical characteristic of a material, $J/(m^3\ C)$

A FUNCTION ESTIMATION problem can be formed if k and C are known and T-measurements are available but $q_0 = q_0(t)$ is an unknown function



**Where are the measurement errors?
How can they be described?**

Eight Standard Statistical Assumptions

1. Additive errors,

$$Y_i = T_i + \varepsilon_i$$

2. Zero mean errors

$$E(\varepsilon_i) = 0$$

3. Constant standard deviation,

$$\sigma_i = \sigma = \text{constant}$$

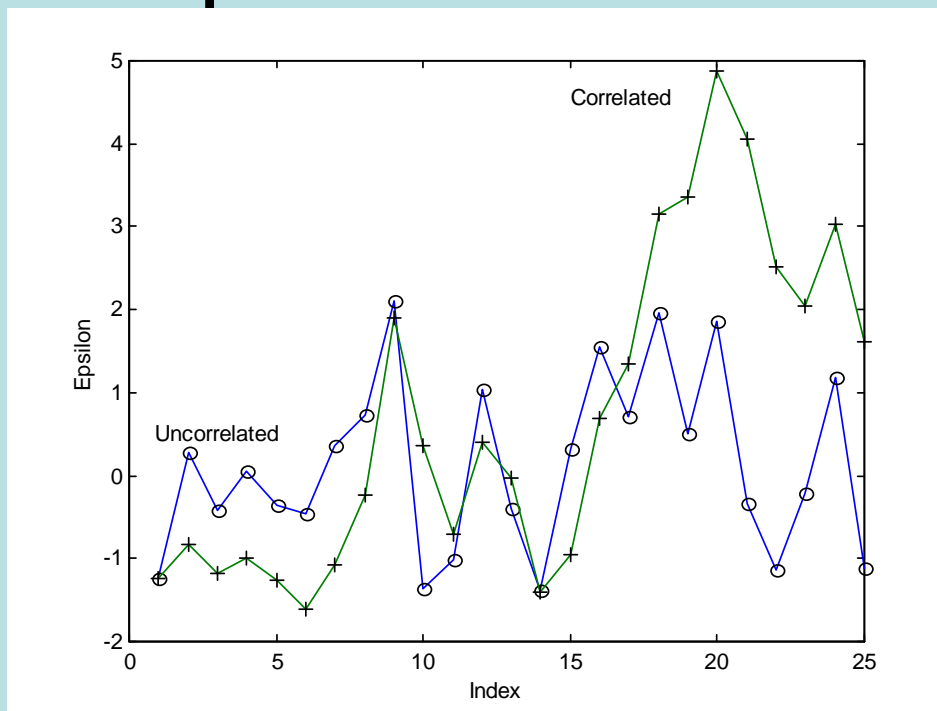
4. Uncorrelated errors

5. Normal distribution of errors

6. Known statistical parameters

7. Errors only in the dependent variable

8. No prior information



Uncorrelated, $\sigma = 1$

13 runs

Correlated, $\sigma = 1$

$\rho = 0.9$, 6 runs

$$\varepsilon_i = \rho\varepsilon_{i-1} + u_i$$

randnum

A PRINCIPLE OF PARAMETER ESTIMATION

We want to design experiments so that the scaled sensitivity coefficients are large and uncorrelated.

i^{th} Sensitivity Coefficient, $\frac{\partial T}{\partial \beta_i}$

i^{th} Scaled Sensitivity Coefficient, $\beta_i \frac{\partial T}{\partial \beta_i}$

Scaled Sensitivity Coefficients are important. Why??

COMPUTATION OF SCALED SENSITIVITY COEFFICIENTS

1. ANALYTICAL DIFFERENTIATION, GOOD BUT ERROR-PRONE
2. FINITE DIFFERENCE, (Intrinsic Verification)

$$\begin{aligned} & \beta_j \frac{\partial T}{\partial \beta_j} \\ & \approx \beta_j \frac{T(x, t, \beta_1, \dots, \beta_j(1 + \delta), \dots, \beta_p) - T(x, t, \beta_1, \dots, \beta_p)}{\beta_j(1 + \delta) - \beta_j} \\ & \approx \frac{T(x, t, \beta_1, \dots, \beta_j(1 + \delta), \dots, \beta_p) - T(x, t, \beta_1, \dots, \beta_p)}{\delta} \end{aligned}$$

$$\delta = 0.0001 \text{ or } 0.00001$$

TAYLOR SERIES ON PARAMETER VALUES

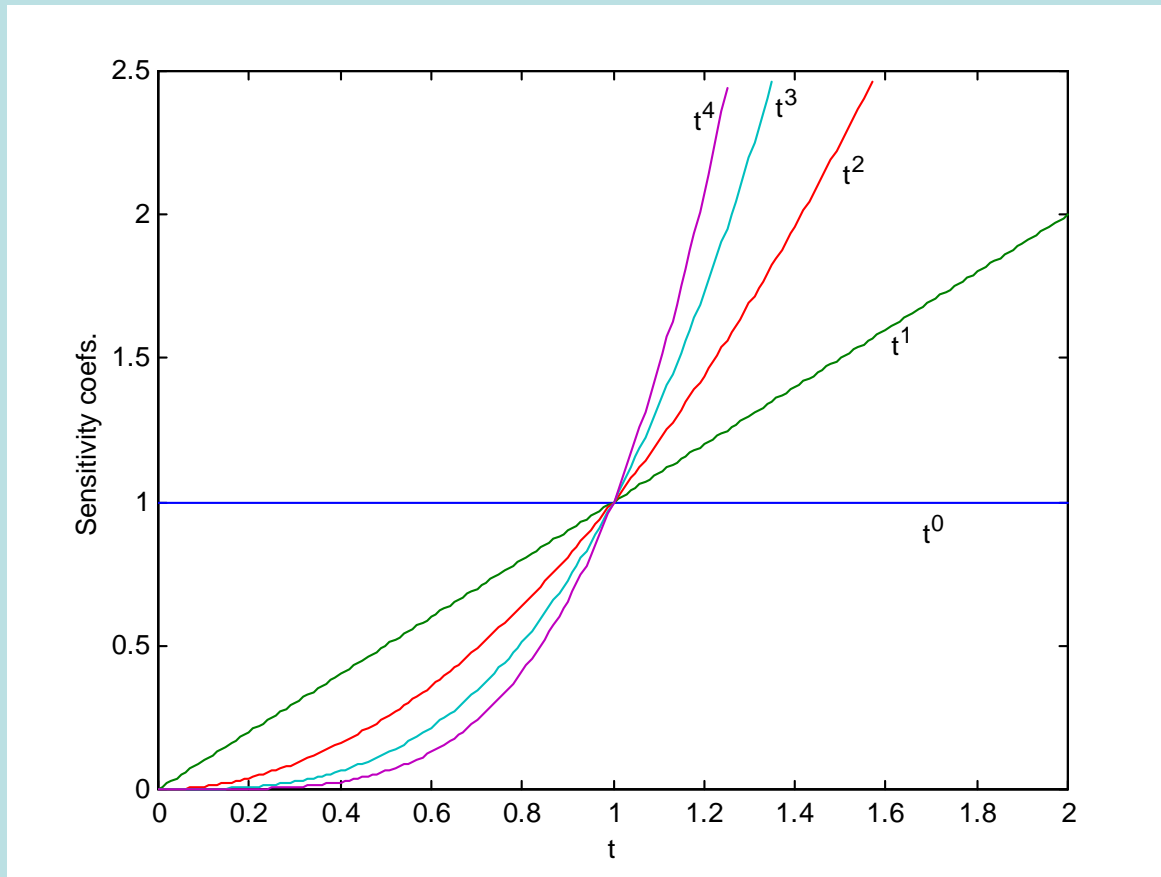
Let $\boldsymbol{\beta}_0 = (\beta_{10}, \beta_{20}, \beta_{30}, \dots, \beta_{p0})$

$$\begin{aligned} T(x, t, \boldsymbol{\beta}) &\approx T(x, t, \boldsymbol{\beta}_0) + \left. \frac{\partial T}{\partial \beta_1} \right|_{\boldsymbol{\beta}_0} (\beta_1 - \beta_{10}) + \dots + \left. \frac{\partial T}{\partial \beta_p} \right|_{\boldsymbol{\beta}_0} (\beta_p - \beta_{p0}) \\ &\approx T(x, t, \boldsymbol{\beta}_0) + \underbrace{\left(\beta_1 \left. \frac{\partial T}{\partial \beta_1} \right|_{\boldsymbol{\beta}_0} \right)}_{\substack{\mathbf{1^{st} \text{ scaled} \\ \text{sensitivity}}} } \frac{\beta_1 - \beta_{10}}{\beta_1} + \dots + \underbrace{\left(\beta_p \left. \frac{\partial T}{\partial \beta_p} \right|_{\boldsymbol{\beta}_0} \right)}_{\substack{\mathbf{p^{th} \text{ scaled} \\ \text{sensitivity}}} } \frac{\beta_p - \beta_{p0}}{\beta_p} \end{aligned}$$

**Want sensitivity coefficients to be “large”
compared to T and uncorrelated with each other**

Polynomial model: $\eta_i = \beta_1 + \beta_2 t_i + \beta_3 t_i^2 + \beta_4 t_i^3 + \beta_5 t_i^4$

$$\frac{\partial \eta_i}{\partial \beta_1} = 1, \quad \frac{\partial \eta_i}{\partial \beta_2} = t_i, \quad \frac{\partial \eta_i}{\partial \beta_3} = t_i^2, \quad \frac{\partial \eta_i}{\partial \beta_4} = t_i^3, \quad \frac{\partial \eta_i}{\partial \beta_5} = t_i^4$$



t^3 and t^4 sens. coef. becoming more correlated.

IDENTIFIABILITY CONDITION:

To estimate parameters, must not have:

$$A_1\beta_1 \frac{\partial T}{\partial \beta_1} + A_2\beta_2 \frac{\partial T}{\partial \beta_2} + \dots + A_p\beta_p \frac{\partial T}{\partial \beta_p} = 0$$

for not all the A_i values equal to zero..

T = temperature/dependent variable

β_i = i^{th} parameter

p = no. of parameters

A_i = dimensionless constant, like 0, +1, -1

If we can show that

$$\beta_1 \frac{\partial T}{\partial \beta_1} + \beta_2 \frac{\partial T}{\partial \beta_2} + \cdots + \beta_p \frac{\partial T}{\partial \beta_p} \neq 0$$

it may be possible to estimate all the parameters simultaneously. Can we show that? Consider the same transient heat conduction example,

$$k \frac{\partial^2 T}{\partial x^2} = C \frac{\partial T}{\partial t}, \quad T(x, 0) = 0$$

$$-k \frac{\partial T}{\partial x}(0, t) = q_0, \quad \frac{\partial T}{\partial x}(L, t) = 0$$

DERIVATION OF SUM OF SCALED SENSITIVITY COEFFICIENTS RELATIONSHIP

Consider the same problem in dimensionless form:

$$\tilde{T} = \frac{T}{q_0 L}, \quad \tilde{x} = \frac{x}{L}, \quad \tilde{t} = \frac{kt}{CL^2}, \quad T = \frac{q_0 L}{k} \tilde{T}(\tilde{x}, \tilde{t}(k, C, L, t))$$

Find k sensitivity coefficient using chain rule

$$\begin{aligned} \left. \frac{\partial T}{\partial k} \right|_{x,t,C \text{ fixed}} &= -\frac{q_0 L}{k^2} \tilde{T} + \frac{q_0 L}{k} \left[\left. \frac{\partial \tilde{T}}{\partial \tilde{t}} \right|_{\tilde{x}} \frac{\partial \tilde{t}}{\partial k} \right] \\ &= -\frac{q_0 L}{k^2} \tilde{T} + \frac{q_0 L}{k} \left[\left. \frac{\partial \tilde{T}}{\partial \tilde{t}} \right|_{\tilde{x}} \frac{t}{CL^2} \right] \end{aligned}$$

Multiply by k to get

$$k \frac{\partial T}{\partial k} \Big|_{x,t,C \text{ fixed}} = -\frac{q_0 L}{k} \tilde{T} + \frac{q_0 L}{k} \tilde{t} \frac{\partial \tilde{T}}{\partial \tilde{t}} \Big|_{\tilde{x}} = -T + \frac{q_0 L}{k} \tilde{t} \frac{\partial \tilde{T}}{\partial \tilde{t}} \Big|_{\tilde{x}}$$

Repeat for C sensitivity coefficient

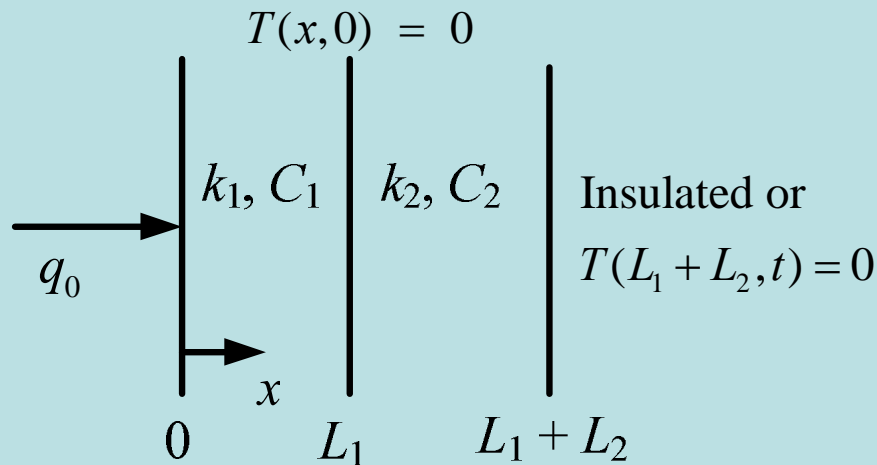
$$\frac{\partial T}{\partial C} \Big|_{x,t,k \text{ fixed}} = \frac{q_0 L}{k} \frac{\partial \tilde{T}}{\partial \tilde{t}} \Big|_{\tilde{x}} \frac{\partial \tilde{t}}{\partial C} = \frac{q_0 L}{k} \frac{\partial \tilde{T}}{\partial \tilde{t}} \Big|_{\tilde{x}} \left(-\frac{kt}{C^2 L^2} \right)$$

$$C \frac{\partial T}{\partial C} \Big|_{x,t,k \text{ fixed}} = -\frac{q_0 L}{k} \tilde{t} \frac{\partial \tilde{T}}{\partial \tilde{t}} \Big|_{\tilde{x}}$$

Add these two scaled sensitivities to get

$$k \frac{\partial T(x,t)}{\partial k} + C \frac{\partial T(x,t)}{\partial C} = -T(x,t)$$

ANOTHER EXAMPLE, TWO PLATES/MATERIALS



We can derive:

$$k_1 \frac{\partial T(x,t)}{\partial k_1} + C_1 \frac{\partial T(x,t)}{\partial C_1} + k_2 \frac{\partial T(x,t)}{\partial k_2} + C_2 \frac{\partial T(x,t)}{\partial C_2} = -T(x,t)$$

For temperature boundary condition at $x = 0$:

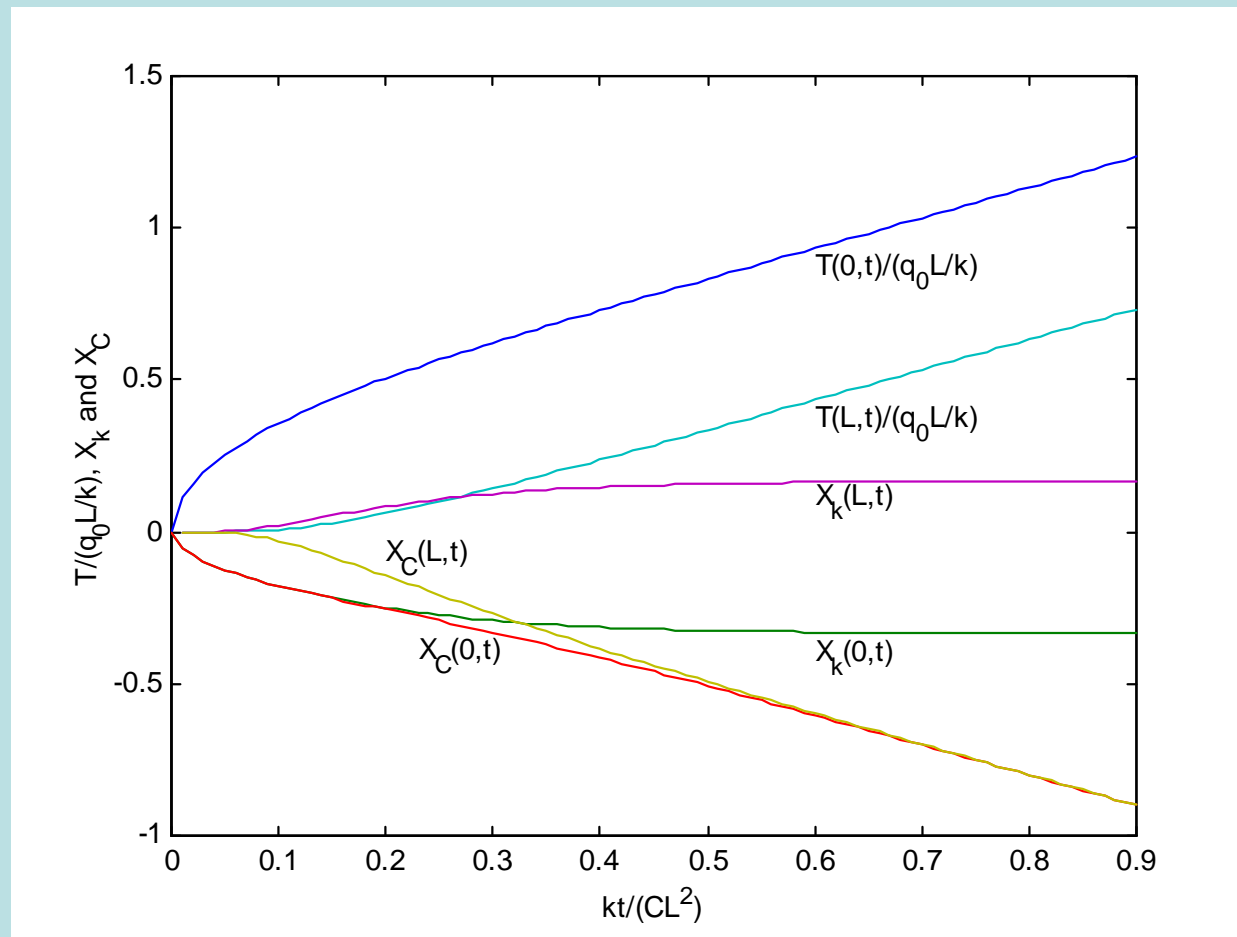
$$k_1 \frac{\partial T(x,t)}{\partial k_1} + C_1 \frac{\partial T(x,t)}{\partial C_1} + k_2 \frac{\partial T(x,t)}{\partial k_2} + C_2 \frac{\partial T(x,t)}{\partial C_2} = 0$$

Insights from Sensitivities Summation Relation

1. If T is not zero, scaled sensitivities might be linearly independent. Then k and C might be simultaneously and independently estimated.
2. For $T > 0$, sum of scaled sensitivities < 0 . Sensitivities tend to be negative for $T > 0$. If each sensitivity < 0 , each sensitivity is less than $\text{abs}(T)$
3. As more parameters are estimated, it becomes more difficult since the sum is still $-T$.

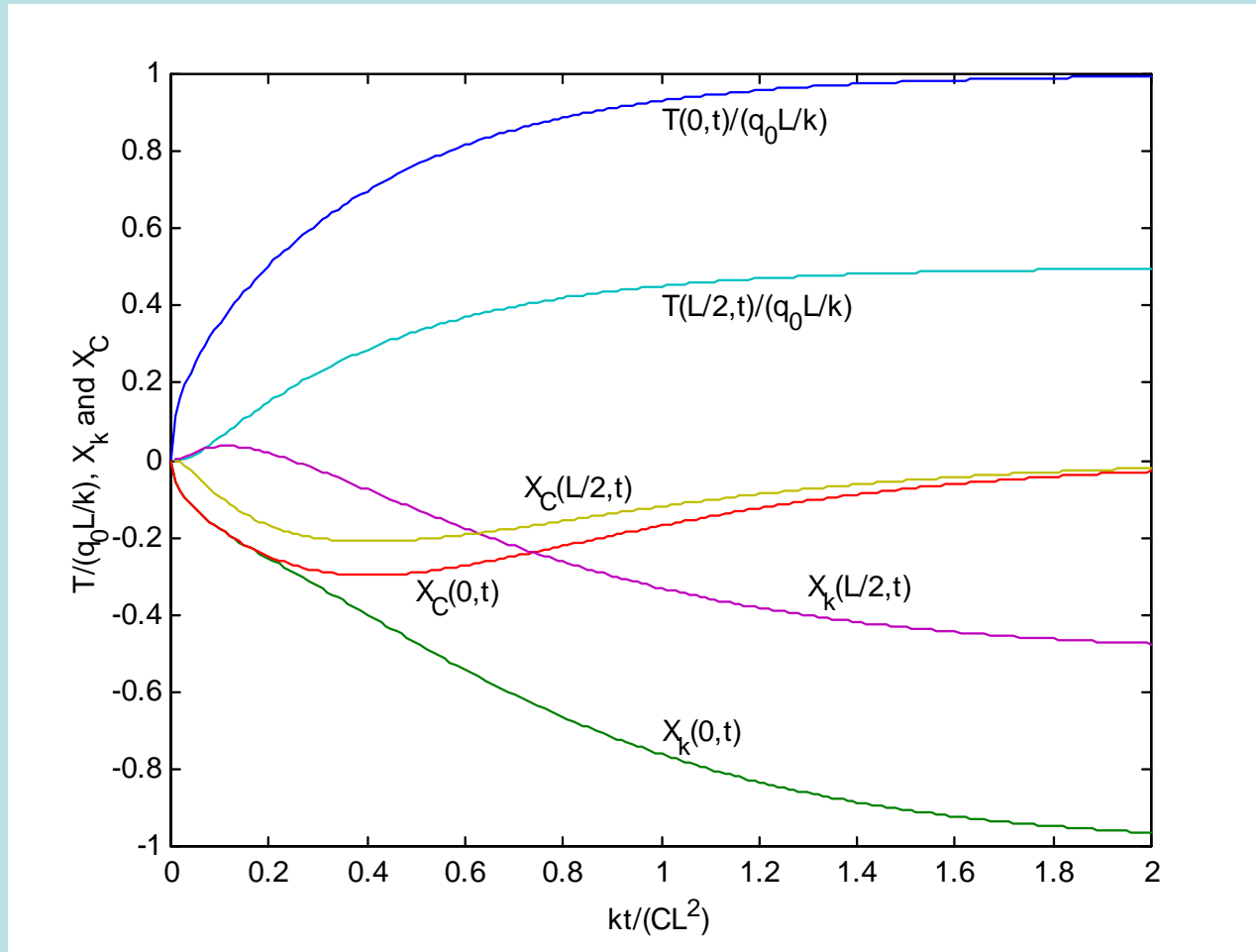
T & scaled sens. coefficients for single material case

$$X_k \equiv \frac{1}{q_0 L / k} k \frac{\partial T}{\partial k}, \quad X_c \equiv \frac{1}{q_0 L / k} C \frac{\partial T}{\partial C}$$



Temperature and scaled sensitivity coefficients for X21B10T0 case (const. heat flux @ $x = 0$, $T(L,t) = 0$)

$$X_k \equiv \frac{1}{q_0 L / k} k \frac{\partial T}{\partial k}, \quad X_c \equiv \frac{1}{q_0 L / k} C \frac{\partial T}{\partial C}$$



OBSERVATIONS – SENS. COEF.

1. Sensitivity coefficients tend to be negative.
2. Coefficients tend to be “large.”
3. Coefficients have different shapes, that is, are uncorrelated.

However, for measurements just at $x = 0$ and until dimensionless time until about 0.2

$$X_k = X_C \text{ or } k \frac{\partial T}{\partial k} - C \frac{\partial T}{\partial C} = 0$$

Then, k and C cannot be estimated.

PARAMETER ESTIMATION PROBLEMS ARE NONLINEAR IF THE SENSITIVITY COEFFICIENTS ARE FUNCTIONS OF THE PARAMETERS

Linear model in the parameters:

$$\eta_i = \beta_1 + \beta_2 t_i$$

$\eta_i =$ Dependent variable

$\beta_1, \beta_2 =$ Parameters

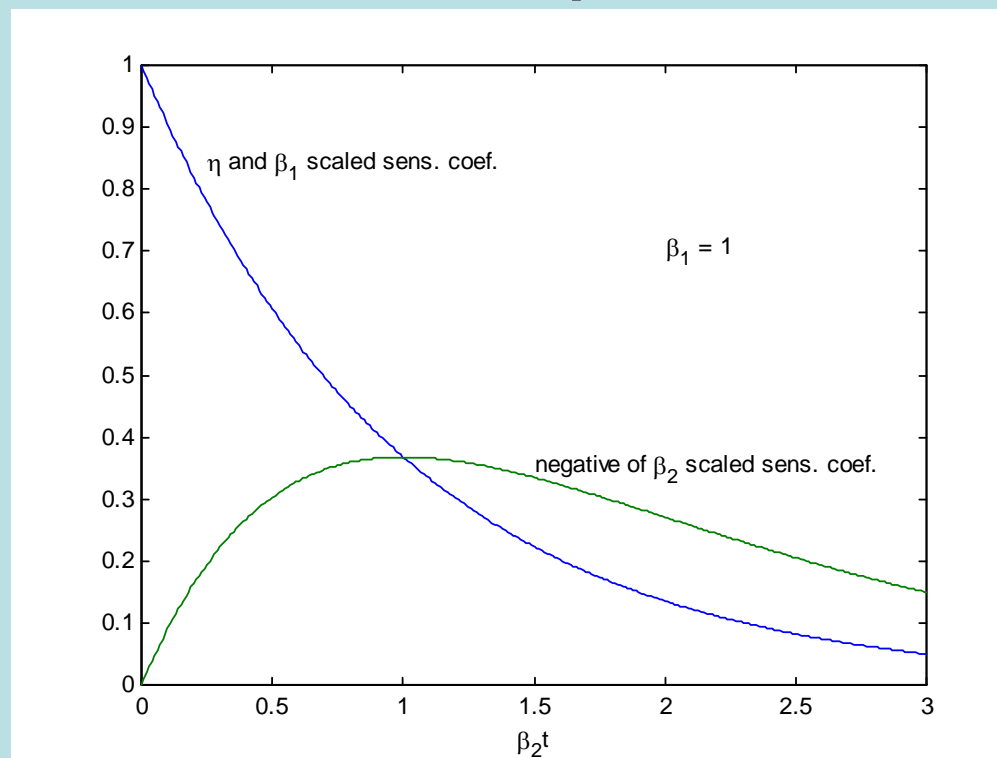
$t_i =$ Independent variable

$\frac{\partial \eta_i}{\partial \beta_1} = 1, \quad \frac{\partial \eta_i}{\partial \beta_2} = t_i,$ **Not functions of parameter, then linear estimation problem**

Nonlinear model in the parameters:

$$\eta_i = \beta_1 e^{-\beta_2 t_i}, \quad \frac{\partial \eta_i}{\partial \beta_1} = e^{-\beta_2 t_i}, \quad \frac{\partial \eta_i}{\partial \beta_2} = -\beta_1 t_i e^{-\beta_2 t_i}$$

Since sens. coef. functions of parameters, the parameter estimation problem is nonlinear..



HEAT CONDUCTION EXAMPLE, k and C parameters

$$k \frac{\partial^2 T}{\partial x^2} = C \frac{\partial T}{\partial t}, \quad -k \frac{\partial T}{\partial x}(0, t) = q_0, \quad \frac{\partial T}{\partial x}(L, t) = 0, \quad T(x, 0) = 0$$

$$T = T(x, t, k, C, q_0, L), \quad \text{Let } X_c \equiv \left. \frac{\partial T}{\partial C} \right|_{k \text{ fixed}}$$

$$k \frac{\partial^2 X_c}{\partial x^2} = \frac{\partial T}{\partial t} + C \frac{\partial X_c}{\partial t}, \quad -k \frac{\partial X_c}{\partial x}(0, t) = 0, \quad \frac{\partial X_c}{\partial x}(L, t) = 0$$

$$X_c(x, 0) = 0$$

NONLINEAR P. E. SINCE X_c IS FUNCTION OF PARAMETERS.

ESTIMATION PROCEDURES,

1. ORDINARY LEAST SQUARES

$$S = \sum_{i=1}^n (Y_i - \eta_i)^2 = (\mathbf{Y} - \boldsymbol{\eta})^T (\mathbf{Y} - \boldsymbol{\eta})$$

$$\mathbf{Y} = \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix} = \text{Meas. vector}, \quad \boldsymbol{\eta} = \begin{bmatrix} \eta_1 \\ \vdots \\ \eta_n \end{bmatrix} = \text{model vector}$$

For a model linear in terms of the parameters

$$\eta_i = \beta_1 X_{i1} + \beta_2 X_{i2} + \cdots + \beta_p X_{ip} \quad \text{or} \quad \boldsymbol{\eta} = \mathbf{X}\boldsymbol{\beta}$$

$$\mathbf{X} = \begin{bmatrix} X_{11} & \cdots & X_{1p} \\ \vdots & \ddots & \vdots \\ X_{n1} & \cdots & X_{np} \end{bmatrix} \quad \text{Sensitivity matrix}$$

$$\boldsymbol{\beta} = \begin{bmatrix} \beta_1 \\ \vdots \\ \beta_p \end{bmatrix} \quad \text{Parameter vector}$$

ORDINARY LINEAR LEAST SQUARES ESTIMATOR

$$\mathbf{b}_{OLS} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{Y}$$

Excellent estimator if standard statistical are valid

**NOT ALL STANDARD STAT. ASSUMPTIONS VALID:
 USE MAXIMUM LIKELIHOOD IF STANDARD DEVIATION IS
 NOT CONSTANT, ERRORS CORRELATED & NORMAL**

Need covariance matrix

$$\Psi = \begin{bmatrix} \sigma_1^2 & \sigma_{12} & \cdots & \sigma_{1n} \\ \sigma_{12} & \sigma_2^2 & \cdots & \sigma_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{1n} & \sigma_{2n} & \cdots & \sigma_n^2 \end{bmatrix}$$

$\sigma_i^2 = \text{variance of } Y_i = \eta_i + \varepsilon_i = V(Y_i) = V(\varepsilon_i)$
 $\sigma_i = \text{standard deviation}$
 $\sigma_{ij} = \text{cov}(Y_i, Y_j)$

$$S_{ML} = (\mathbf{Y} - \boldsymbol{\eta})^T \boldsymbol{\Psi}^{-1} (\mathbf{Y} - \boldsymbol{\eta})$$

For Linear Model: $\mathbf{b}_{ML} = (\mathbf{X}^T \boldsymbol{\Psi}^{-1} \mathbf{X})^{-1} \mathbf{X}^T \boldsymbol{\Psi}^{-1} \mathbf{Y}$

IF PRIOR INFORMATION IS AVAILABLE, USE MAXIMUM A POSTERIORI ESTIMATION

$\boldsymbol{\mu}_\beta$ = prior parameter values

\mathbf{V}_β = covariance matrix of the prior parameters

For linear-in-parameters model,

$$\mathbf{b}_{MAP} = \boldsymbol{\mu}_\beta + (\mathbf{X}^T \boldsymbol{\Psi}^{-1} \mathbf{X} + \mathbf{V}_\beta^{-1})^{-1} \mathbf{X}^T \boldsymbol{\Psi}^{-1} (\mathbf{Y} - \mathbf{X} \boldsymbol{\mu}_\beta)$$

We have given estimators for linear-in-parameter models. Usually non-linear, however.

NONLINEAR-IN-PARAMETERS ESTIMATION

One way is Gauss linearization.

$$S = (\mathbf{Y} - \boldsymbol{\eta}(\boldsymbol{\beta}))^T \mathbf{W} (\mathbf{Y} - \boldsymbol{\eta}(\boldsymbol{\beta})) + (\boldsymbol{\mu}_\beta - \boldsymbol{\beta})^T \mathbf{V}_\beta^{-1} (\boldsymbol{\mu}_\beta - \boldsymbol{\beta})$$

\mathbf{W} = Weighting matrix

$$\mathbf{b}^{(k+1)} = \mathbf{b}^{(k)} + \mathbf{P}^{(k)} \left[\mathbf{X}^{T(k)} \mathbf{W} (\mathbf{Y} - \boldsymbol{\eta}^{(k)}) \right] + \mathbf{V}_\beta^{-1} (\boldsymbol{\mu}_\beta - \mathbf{b}^{(k)})$$

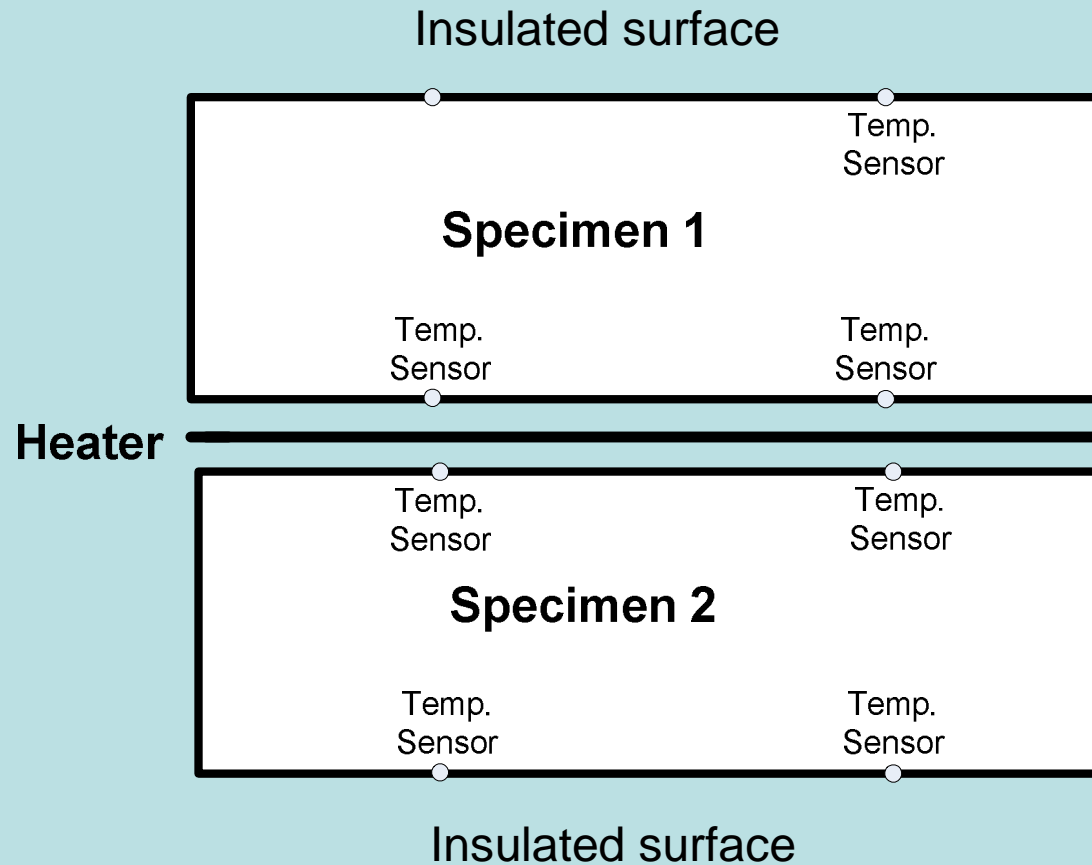
$$\mathbf{P}^{-1(k)} = \mathbf{X}^{T(k)} \mathbf{W} \mathbf{X}^{(k)} + \mathbf{V}_\beta^{-1}$$

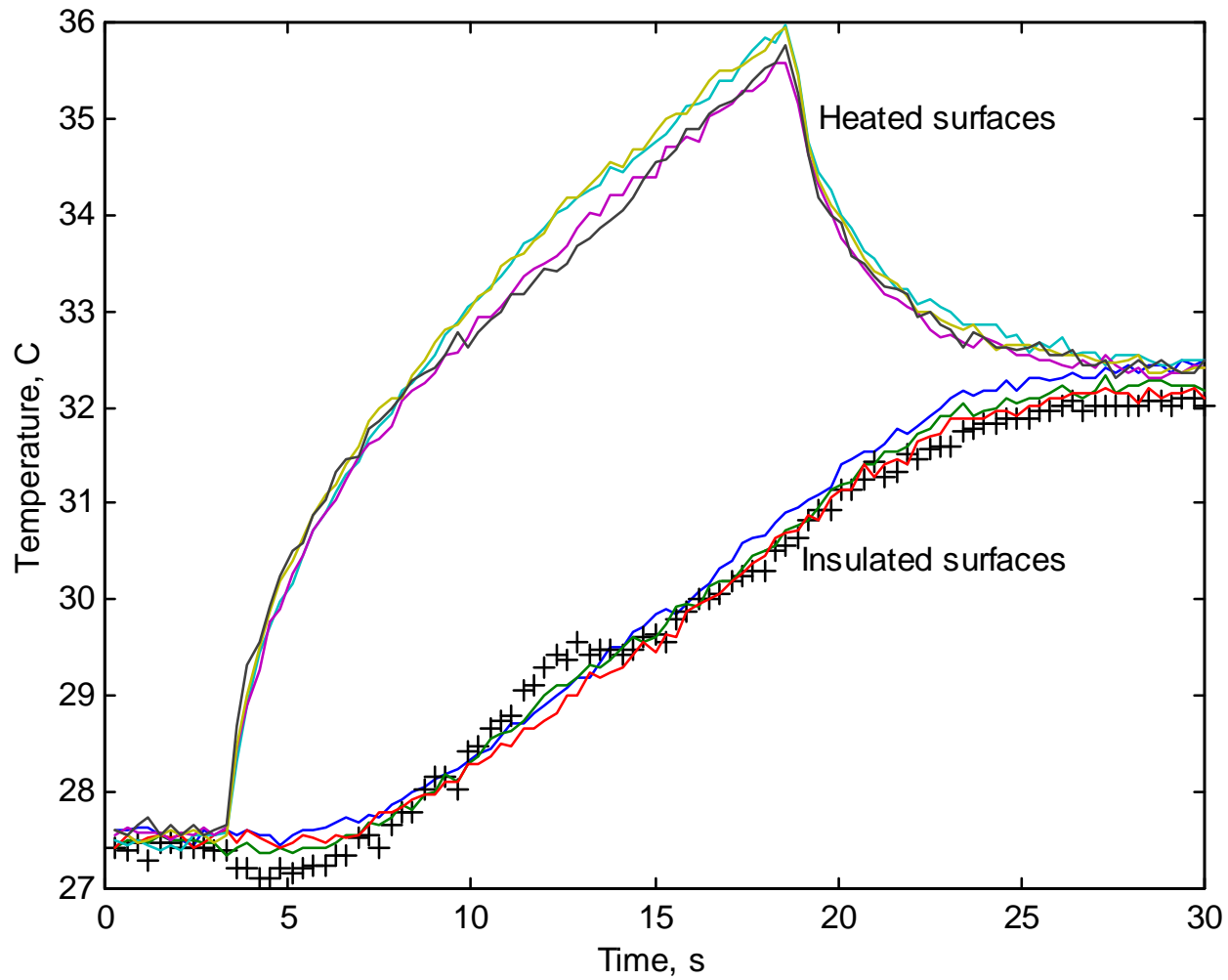
MODIFICATIONS TO GAUSS LINEARIZATION

1. Use prior information (maybe non-informative prior)
2. Modify step size (Box-Kanemasu)
3. Do sequentially, add one measurement at a time

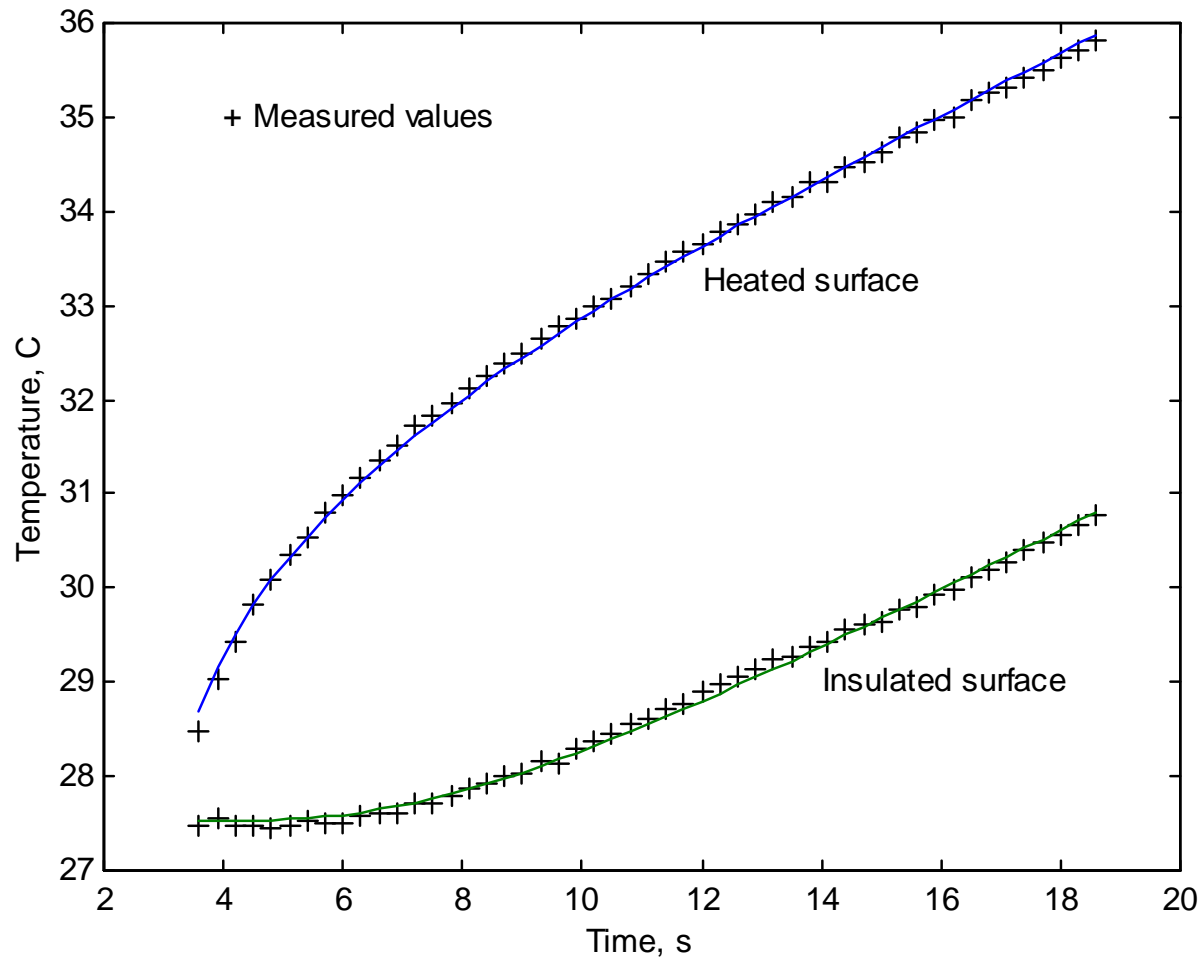
Transient Heat Conduction Experiment to find k and C

Specimens 0.0254 m thick, 0.076 m diameter, $q_0 = 30,300 \text{ W/m}^2$





Measured T_s . Curve with + sign inaccurate. Drop its data.



Averaged measured & computed T values. Good agreement.

Estimated parameter values of Armco iron

Using PROP1D

Thermal conductivity, 75.4 +/- 1.2 W/m-C, +/-1.6%

Volumetric heat capacity, 3,700,000 +/-56,000 J/m³-C, +/-1.5%

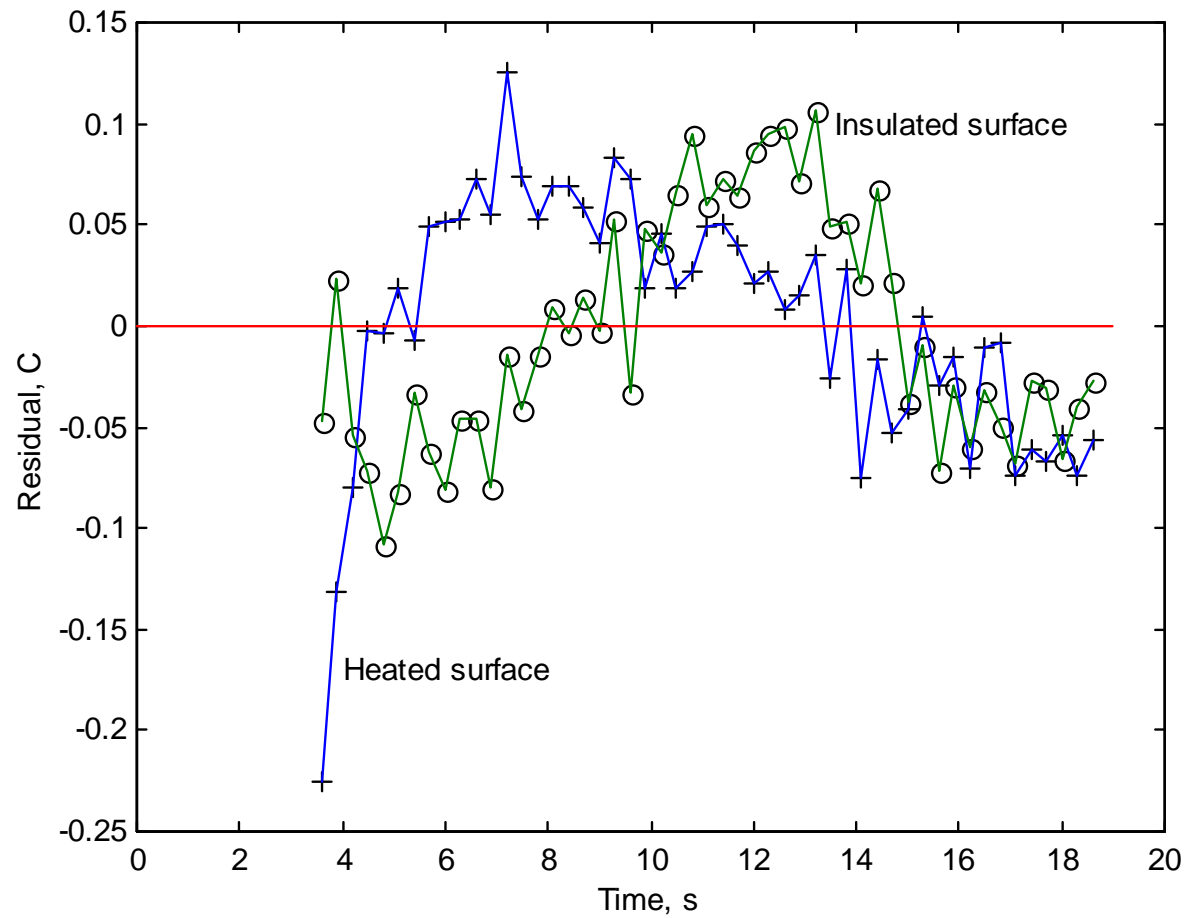
RMS, 0.0599 C

Incropera and DeWitt 99.75 % pure Armco iron at 27 C

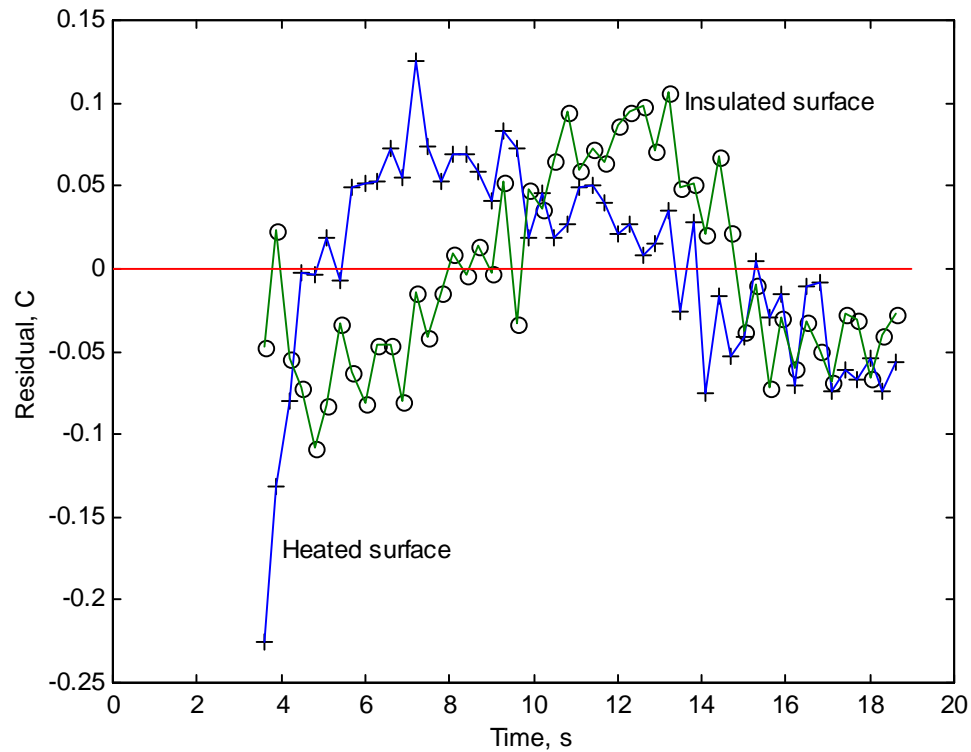
Thermal conductivity 72.7 W/m-C, within 3.6%

Volumetric heat capacity, 3,520,000 J/m³-C, within 4.9%

ITER	RMS	TH COND	VOL HEAT
0		50	.25E+7
1	2.092	66.83	.3314E+07
2	0.522	74.39	.3664E+07
3	0.079	75.35	.3706E+07
4	0.060	75.36	.3706E+07

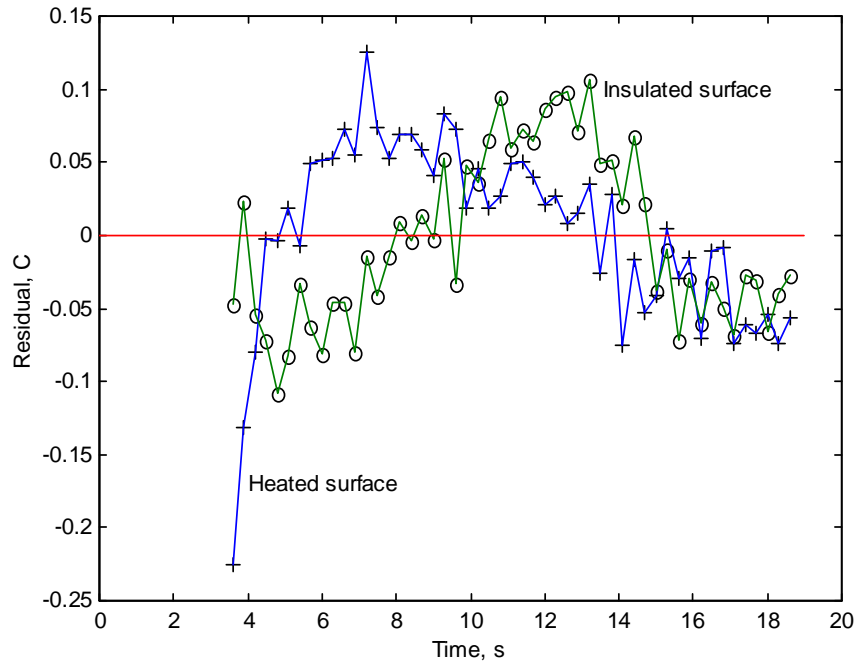


Residuals. What can we learn? Compare with std. assumptions



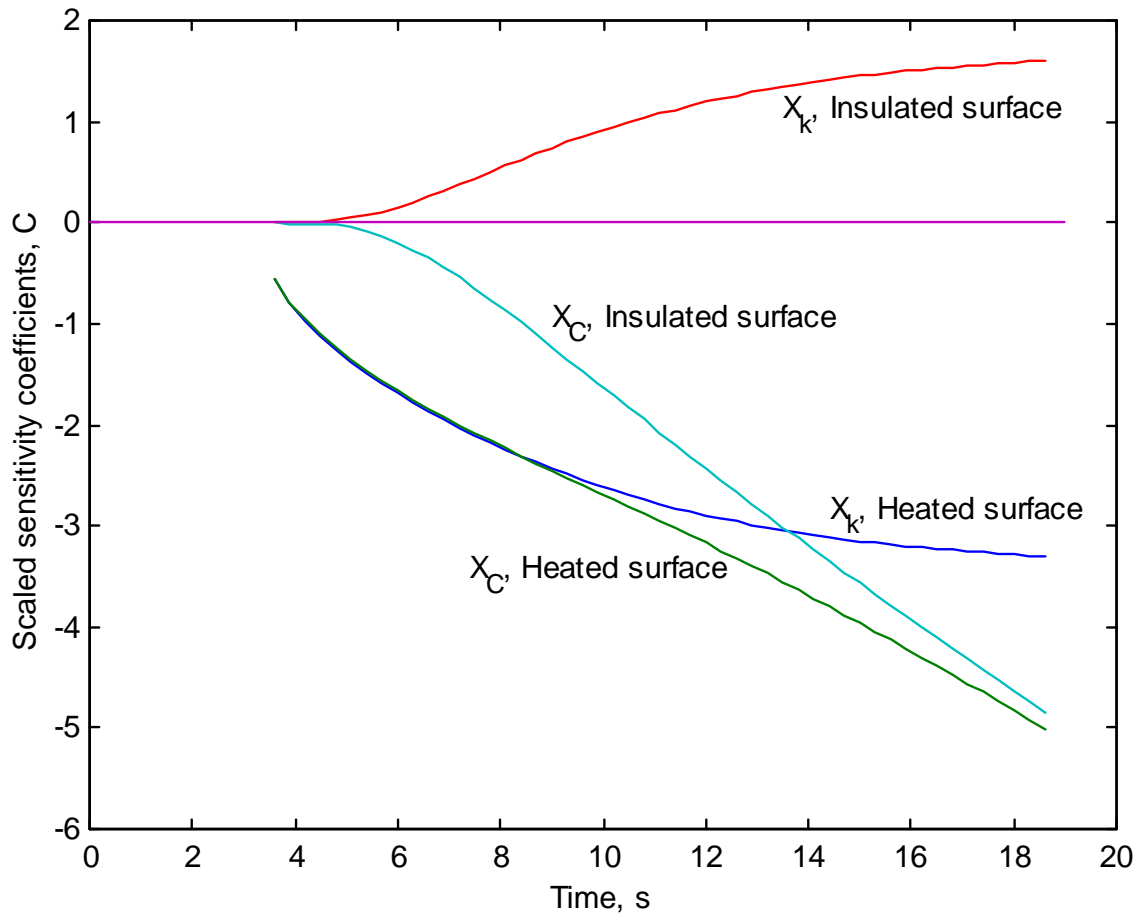
Comparison with standard statistical assumptions

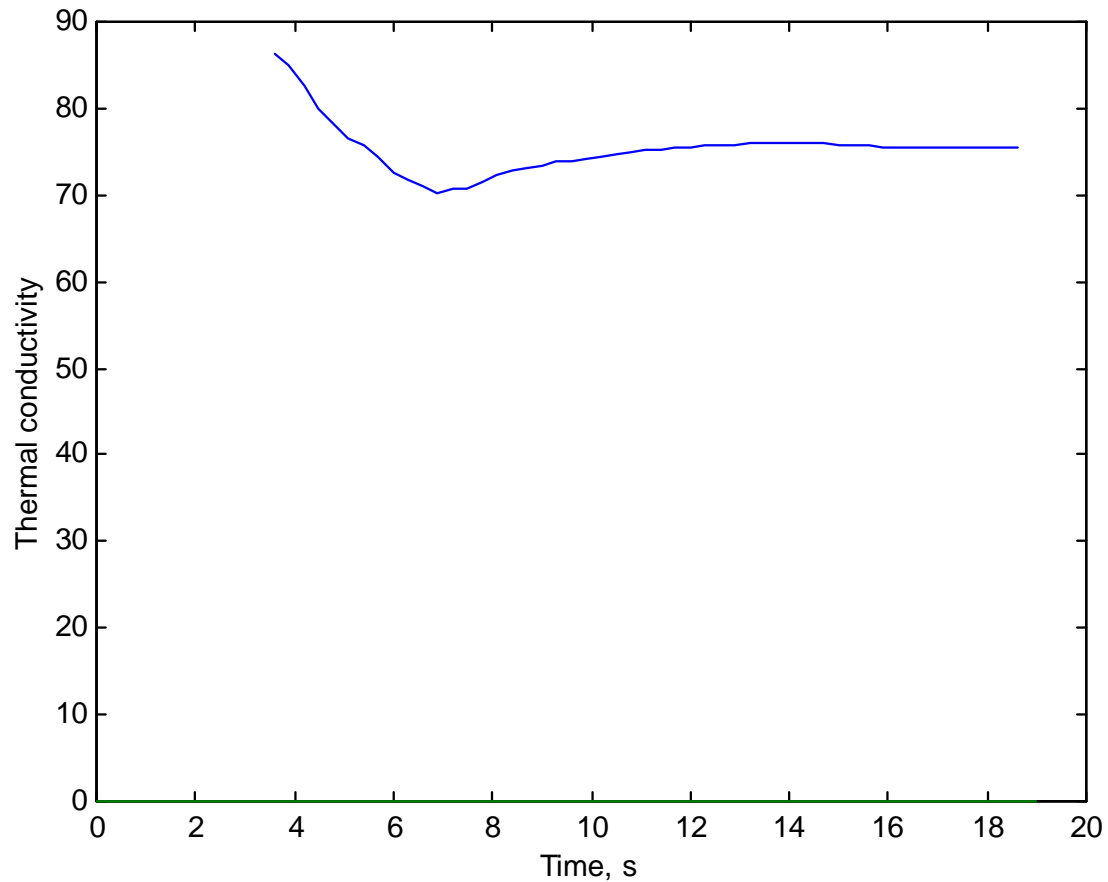
1. Additive errors. O.K.
2. Zero mean errors. O.K.
3. Constant variance. O.K.
4. Non-correlated errors. Not valid
5. Gaussian errors. O.K.



Significant features:

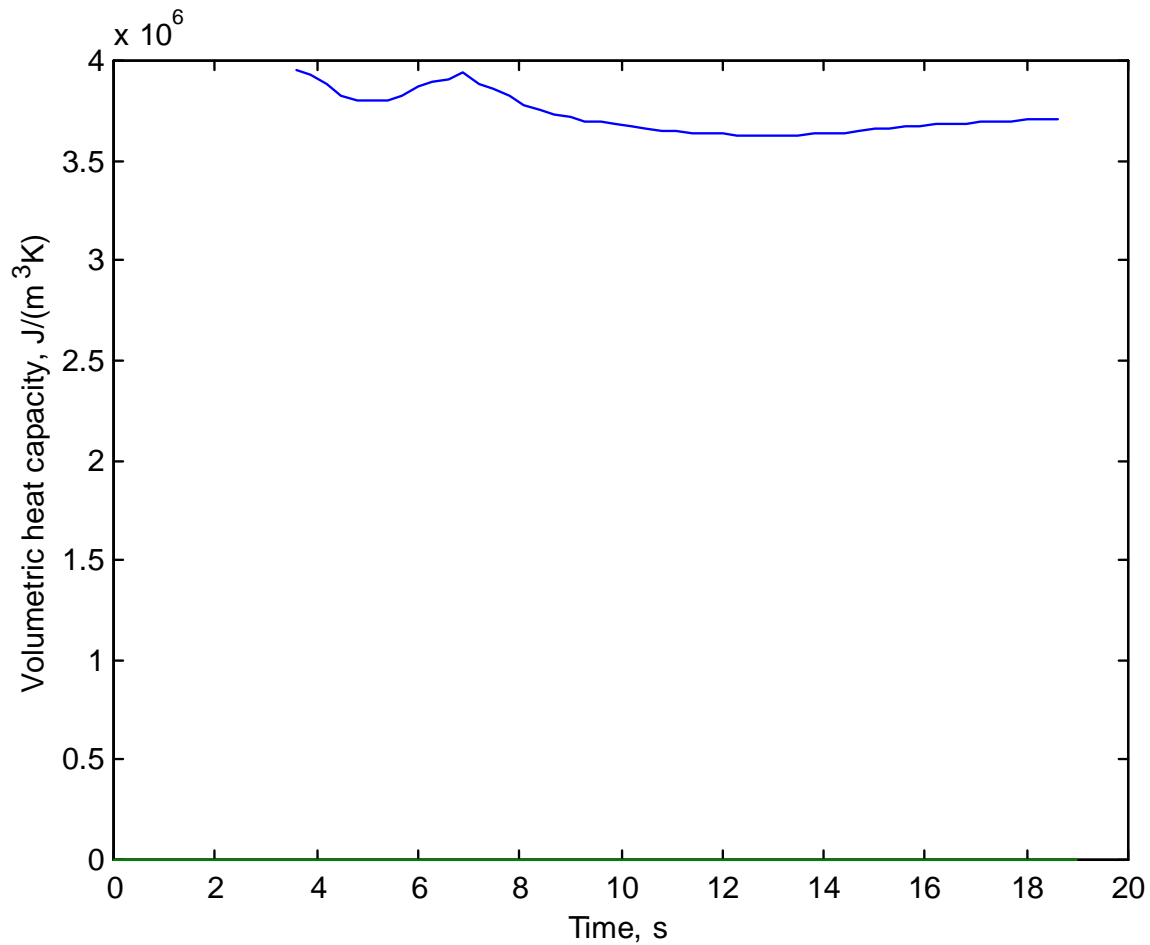
1. Maximum magnitude residual is at 1st time at heated surf.
 Heater not modeled. Starting time may not be precisely known
2. Characteristic “signature” of both $x = 0$ and L residuals are negative at final times. Drift lower. Heat loss not modeled?





Time	k
9.9	74.20
12.0	75.54
14.1	75.91
16.2	75.55
18.6	75.36
2.3%	

SEQUENTIAL k ESTIMATES. Sensitivity coefficients evaluated at converged k & C values.



Time C x 10⁻⁶

9.9 3.67

12.0 3.63

14.1 3.64

16.2 3.68

18.6 3.71

2.2%

SEQUENTIAL C ESTIMATES.

TWO MORE IMPORTANT TOPICS:

1. Confidence regions
2. Optimal experiment design

Both use sensitivity coefficients.

Consider briefly **OPTIMAL EXPERIMENTS**

Let $\mathbf{X} = \begin{bmatrix} X_{11} & \cdots & X_{1p} \\ \vdots & \ddots & \vdots \\ X_{n1} & \cdots & X_{np} \end{bmatrix}$ **= sensitivity matrix**

Criterion: Maximize $|\mathbf{X}^T \mathbf{X}|$ Subject to constraints

Possible constraints:

- 1. Given No. of measurement points**
- 2. Given range of dependent variable**
- 3. Cost or whatever**

Adjustable conditions

- 1. Location and no. of sensors**
- 2. Duration of experiment**
- 3. Type of boundary conditions**
- 4. Time variation of boundary conditions**

CONCLUSIONS

1. It is important learn about and characterize the measurement errors. Should check assumptions using residuals.
2. The scaled sensitivity coefficients can give important insight into the design of experiments. Also aid in determining which parameters can be estimated.
3. Do not over-parameterize! Adding parameters makes estimation more difficult
4. Some actual experiments are described and related sensitivity coefficients given. Sequential analysis can give insight into results.
5. The residuals are important. Examine to determine if any characteristic signatures are present. Should model be improved?